

23. REPEATED MEASURES TYPES OF DESIGNS

23.1 INTRODUCTION

We treat repeated measures types of designs as a unique topic since in psychology, social work, education...etc. these types of designs are extremely popular. These designs can be viewed as having two components. They can be considered as examples of nested designs with specific assumptions on the correlations between observations. We will briefly discuss both points of view. Psychologists, when conducting studies that use humans and animals, commonly take observations on the same person/animal a number of times, say $j = 1, 2$, psychology when studying how learning occurs over time. Note that in contrast to the randomized block design, treatments are not randomly allocated to blocks.

Subject	Time			
	1	2	. . .	t
S_1	y_{11}	y_{12}	. . .	y_{1t}
S_2	t_{21}	y_{22}	. . .	y_{2t}
.	.			.
.	.			.
.	.			.
S_n	y_{n1}	y_{n2}	. . .	y_{nt}

23.2 NESTED OR UNIVARIATE APPROACH

The above design is very similar to the two factor factorial model with one observation per treatment. The effect of time is seen within people (subjects).

The EMS table would look like the following.

Source	DF	n	t	1	EMS
		R	F	R	
		i	j	k	
S_i	n-1	1	t	1	$\sigma_e^2 + t\sigma_S^2$
T_j	t-1	n	0	1	$\sigma_e^2 + \sigma_{ST}^2 + n\sigma_T^2$
ST_{ij}	(n-1)(t-1)	1	0	1	$\sigma_e^2 + \sigma_{ST}^2$
$E_{k[ij]}$	0	1	1	1	σ_e^2 (not available)

Since there is only one observation per cell, the error term, e , is not estimable. Treating the ST interaction term as error is the proper F test for the Time (T) factor.

The ANOVA table is usually reported as follows.

Source	DF	MS	F
Between Subjects (S_i)	n-1		
Within Subjects			
T_j	t-1	MST	$\frac{MST}{MSE}$
Error	(n-1)(t-1)	MSE	
Total	nt-1		

The proper F test of the Time factor is $\frac{MST}{MSE}$.

23.3 CORRELATION ASSUMPTIONS

As mentioned previously given that *between* person variability is greater than *within* person variability, it is more efficient for the psychologist to block on people.

The ANOVA model for this design includes an effect due to Time and Subjects as mentioned previously, however, the assumption of is stressed. In other words the correlation of y_{ij} is constant and takes the form

$$\begin{vmatrix} 1 & \rho & . & . & . & . & \rho \\ \rho & 1 & & & & & . \\ . & & . & & & & . \\ . & & & . & & & . \\ . & & & & . & & . \\ . & & & & & . & . \\ \rho & . & . & . & . & . & 1 \end{vmatrix}$$

The symmetry condition can be assessed using the Huynh and Feldt (1970) procedure. For example, if observations taken closely in time were more highly correlated than those taken further apart, the compound symmetry condition would be violated. SPSS gives a test of significance called the Mauchly test of sphericity and corresponding epsilon multipliers that correct for violations in this assumption. Note the inclusion of the *compound symmetry* assumption in addition to the other linear model assumptions, ie normality, etc. Sometimes a different **MULTIVARIATE APPROACH** is taken to this same problem. Consider the t times as t variables recorded on each person. With this approach one assumes the variables are distributed according to a multivariate normal distribution and there is no assumption of symmetry or circularity. Unfortunately for the student both the univariate and multivariate material are printed in spss and they are mixed in the printout.

23.4 EXAMPLE

Siegel, P.S., Activity level as a function of physically forced inaction, *J. of Psychol*, 1946, 285-291 examined how activity level in rats was a function of confinement. An animal was confined for 0, 6, 12 and 24 hours and activity recorded after the period of confinement. The data is given below.

Rat	0 Hours	6 Hours	12 Hours	24 Hours
1	232	244	213	272
2	216	212	191	269
3	112	62	69	119
4	219	119	200	251
5	292	165	187	287
6	179	106	189	217
7	264	271	354	365
8	247	260	295	305
9	259	241	196	211
10	195	118	150	184
11	140	121	102	136
12	244	189	229	240
13	364	326	329	303
14	302	282	292	362
15	312	233	225	306
16	350	283	312	430
Mean	245	202	221	266

[Click here for the SPSS windows analysis of this problem](#)

Syntax Computer Implementation

```
data list      /rat 1-3, time 4-5, y 6-9
begin data
01 1 232
01 2 244
01 3 213
01 4 272
02 1 216
02 2 212
02 3 191
02 4 269
03 1 112
03 2 062
03 3 069
03 4 119
04 1 219
04 2 119
04 3 200
```

04 4 251
05 1 292
05 2 165
05 3 187
05 4 287
06 1 179
06 2 106
06 3 189
06 4 217
07 1 264
07 2 271
07 3 354
07 4 365
08 1 247
08 2 260
08 3 295
08 4 305
09 1 259
09 2 241
09 3 196
09 4 211
10 1 195
10 2 118
10 3 150
10 4 184
11 1 140
11 2 121
11 3 102
11 4 136
12 1 244
12 2 189
12 3 229
12 4 240
13 1 364
13 2 326
13 3 329
13 4 303
14 1 302
14 2 282
14 3 292
14 4 362
15 1 312
15 2 233
15 3 225
15 4 306
16 1 350
16 2 283
16 3 312
16 4 430
end data
frequency general=all/

```

statistics all
manova y by rat(1,16) time(1,4)/
print=cellinfo(means)/
omeans=TABLES(time)/
residuals=casewise plot/
design=rat, time/
finish

```

Tables of means are given below.
Cell Means and Standard Deviations
Variable .. Y

FACTOR	CODE	Mean	Std. Dev.	N
RAT	1			
TIME	1	232.000	.000	1
TIME	2	244.000	.000	1
TIME	3	213.000	.000	1
TIME	4	272.000	.000	1
RAT	2			
TIME	1	216.000	.000	1
TIME	2	212.000	.000	1
TIME	3	191.000	.000	1
TIME	4	269.000	.000	1
RAT	3			
TIME	1	112.000	.000	1
TIME	2	62.000	.000	1
TIME	3	69.000	.000	1
TIME	4	119.000	.000	1
RAT	4			
TIME	1	219.000	.000	1
TIME	2	119.000	.000	1
TIME	3	200.000	.000	1
TIME	4	251.000	.000	1
RAT	5			
TIME	1	292.000	.000	1
TIME	2	165.000	.000	1
TIME	3	187.000	.000	1
TIME	4	287.000	.000	1
RAT	6			
TIME	1	179.000	.000	1
TIME	2	106.000	.000	1
TIME	3	189.000	.000	1
TIME	4	217.000	.000	1
RAT	7			
TIME	1	264.000	.000	1
TIME	2	271.000	.000	1
TIME	3	354.000	.000	1
TIME	4	365.000	.000	1
RAT	8			
TIME	1	247.000	.000	1
TIME	2	260.000	.000	1
TIME	3	295.000	.000	1
TIME	4	305.000	.000	1
RAT	9			
TIME	1	259.000	.000	1
TIME	2	241.000	.000	1

TIME	3	196.000	.000	1
TIME	4	211.000	.000	1
RAT	10			
TIME	1	195.000	.000	1
TIME	2	118.000	.000	1
TIME	3	150.000	.000	1
TIME	4	184.000	.000	1
RAT	11			
TIME	1	140.000	.000	1
TIME	2	121.000	.000	1
TIME	3	102.000	.000	1
TIME	4	136.000	.000	1
RAT	12			
TIME	1	244.000	.000	1
TIME	2	189.000	.000	1
TIME	3	229.000	.000	1
TIME	4	240.000	.000	1
RAT	13			
TIME	1	364.000	.000	1
TIME	2	326.000	.000	1
TIME	3	329.000	.000	1
TIME	4	303.000	.000	1
RAT	14			
TIME	1	302.000	.000	1
TIME	2	282.000	.000	1
TIME	3	292.000	.000	1
TIME	4	362.000	.000	1
RAT	15			
TIME	1	312.000	.000	1
TIME	2	233.000	.000	1
TIME	3	225.000	.000	1
TIME	4	306.000	.000	1
RAT	16			
TIME	1	350.000	.000	1
TIME	2	283.000	.000	1
TIME	3	312.000	.000	1
TIME	4	430.000	.000	1
For entire sample		233.578	79.633	64

Combined Observed Means for TIME

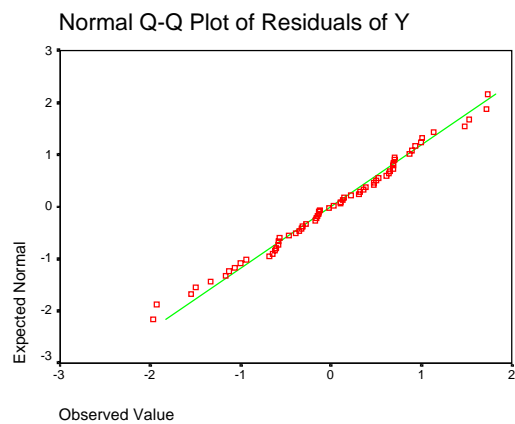
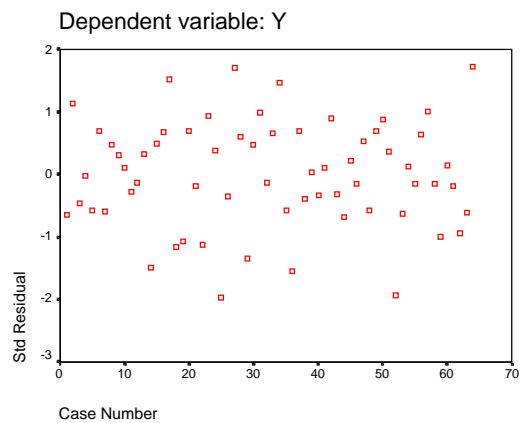
Variable .. Y

TIME			
1	WGT.	245.43750	
	UNWGT.	245.43750	
2	WGT.	202.00000	
	UNWGT.	202.00000	
3	WGT.	220.81250	
	UNWGT.	220.81250	
4	WGT.	266.06250	
	UNWGT.	266.06250	

The coarse ANOVA table is given below. The TIME factor is significant with $F(3,45) = 13.01$, $p < .001$. Planned contrasts or multiple comparisons could be used to further examine the means.

Tests of Significance for Y using UNIQUE sums of squares					
Source of Variation	SS	DF	MS	F	Sig of F
RESIDUAL	43463.45	45	965.85		
RAT	318347.86	15	21223.19	21.97	.000
TIME	37696.30	3	12565.43	13.01	.000
(Model)	356044.16	18	19780.23	20.48	.000
(Total)	399507.61	63	6341.39		

The residuals look reasonable in both plots.



Another Computer Implementation

Another type of implementation is given below that may prove helpful, note that the data is entered differently.

```
data list/ rat 1-2 h0 4-6 h6 8-10 h12 14-16 h24 18-20
begin data
01 232 244 213 272
02 216 212 291 269
. . . . .
. . . . .
```

```

16 350 202 221 266
end data
manova h0 h6 h12 h24
      wsfactors=time(4)/
      wsdesign=time/
      print=cellinfo(means)/
      transform/
      signif(univ,averf)/
      analysis(repeated)/
      design/
finish

```

23.5 MORE COMPLEX MODELS

Suppose an investigator wished to compare both a *between* subject and a *within* subject factor. This type of mixed model is very popular in psychology and social work. The within subject factor is usually a repeated measures time factor (T) which could be the pre, post and post-post test scores of an individual. The between subject factor is usually a factor which represents the type of experimental condition that each person undergoes, call this the Groups factor (G). Subjects are nested within groups ($S_{j(i)}$).

The EMS table would look like

Source	DF	G F i	n R j	t F k	m R l	EMS
G_i	(g-1)	0	n	t	1	$\sigma_e^2 + t \sigma_S^2 + tn\theta_G$
$S_{j(i)}$	g(n-1)	1	1	t	1	$\sigma_e^2 + t \sigma_S^2$
T_k	t-1	g	n	0	1	$\sigma_e^2 + \sigma_{TS}^2 + tg\theta_T$
GT_{ik}	(g-1)(t-1)	0	n	0	1	$\sigma_e^2 + \sigma_{TS}^2 + n\theta_{GT}$
$TS_{kj(i)}$ (ErrorW)	(t-1)g(n-1)	1	1	0	1	$\sigma_e^2 + \sigma_{TS}^2$
$e_l [ijk]$	0	1	1	1	1	σ_e^2 (not available)

The ANOVA table is given below.

TABLE 1

Source	DF	MS	F
Between Subjects (B)	gn-1		
G_i	g-1	MSG	$\frac{MSG}{MSB}$

$S_{j(i)}$ (Error B)	$g(n-1)$	MSB	
Within Subjects (W)			
T_k	$t-1$	MST	$\frac{MST}{MSW}$
GT_{ik}	$(g-1)(t-1)$	MSGT	$\frac{MSGT}{MSW}$
$TS_{kj(i)}$	$(t-1)g(n-1)$	MSW	
Total	$gnt-1$		

There are two error terms, one for the between subject factor G, call this error B, and a within subject error term, error W for the within terms T and TG. The EMS table clearly indicates the proper terms for testing. The above method of breaking down the model is thoroughly reported in Winer and is an example of the nested approach of Section 20.2.

Extensions of this approach are straightforward. For example, consider two fixed factors between and one within repeated.

The ANOVA breakdown would read as follows.

TABLE 2

Source	DF
Between Subjects	
A_i	$a-1$
B_j	$b-1$
AB_{ij}	$(a-1)(b-1)$
$S_{k(ij)}$ (Error B)	$ab(n-1)$
Within Subjects	
C_l	$(c-1)$
AC_{il}	$(a-1)(c-1)$
BC_{ji}	$(b-1)(c-1)$
ABC_{ijl}	$(a-1)(b-1)(c-1)$
$CS_{lk(ij)}$ (Error W)	$(c-1)ab(n-1)$
Total	$abnc-1$

The pattern of analysis with two error terms is repeated in the above example. The A, B and AB terms are tested with the between subject error term whereas the C, AC, BC and ABC terms are tested with the within error term.

23.6 CORRELATION ASSUMPTIONS FOR UNIVARIATE APPROACH

As in the one factor repeated measures design, the T factor of more complex designs assumes the compound symmetry conditions previously discussed. Let us first examine the one factor between and one within design given in Table 1. Referring to Table 1, the F test of the between factor G is not affected. If the Huynh-Feldt conditions do not hold then Greenhouse and Geisser (1959) and Huynh and Feldt (1970) suggest *adjusted* F values be used to determine significance. As we saw previously a quantity called *epsilon* (ϵ) is used to adjust the degrees of freedom of the test. For example to test the effect of the T factor of Table 1 we would compute as before

$$F = \frac{MST}{MSW}$$

however the degrees of freedom for this test are

df numerator = $\epsilon (t - 1)$

df denominator = $\epsilon (t - 1)g (n - 1)$

Note that values of ϵ closer to 0 indicate serious departure from the assumption whereas values closer to 1 indicate less serious violations. When $\epsilon = 1$ we have the original F test. Epsilon ranges between 0 and 1. The MULTIVARIATE approach does not require the circularity/symmetry assumption and as we have seen can readily be applied to the mixed model. SPSS once again mixes the relevant output for both approaches. Remember the different approaches, univariate and multivariate only refer to the within subject variable. The between subject variable is reported the same way for both approaches.

[Click here for the SPSS windows method of analysis of the example below.](#)

23.7 Example

Dunck, E.R., Learning with Secondary Reinforcement Under Two Different Strengths of the Relevant Drive, 1949, as reported in Lindquist (1953), examined how two groups of animals ($n=10$ in each group), one hungry and the other satiated learned a maze task over 20 days. The repeated day factor was collapsed into 4 times. T₁ (Days 1-5), T₂ (6-10), T₃ (11-15) and T₄ (16-20). The data are given below.

Trial Categories

		T_1	T_2	T_3	T_4
Hungry (H_1)	Animal	Days 1-5	Days 6-10	Days 11-15	Days 16-20
	1	3	3	3	3
	2	2	2	4	4
	3	4	2	5	5
	4	3	5	5	5
	5	2	5	5	4
	6	3	5	4	0
	7	1	3	4	1
	8	4	5	3	3
	9	3	5	4	2
	10	0	2	1	0
Satiated (H_2)	11	0	1	2	2
	12	3	0	2	2
	13	2	3	4	4
	14	0	1	2	1
	15	1	3	5	4
	16	1	2	3	2
	17	2	5	1	1
	18	2	3	4	2
	19	3	1	2	1
	20	3	2	2	3

Computer Implementation

```

data list/ HG 1 anim 3-4 T1 6 T2 8 T3 10 T4 12
begin data
1 1 3 3 3 3
1 2 2 2 4 4
.....
.....
.....
2 20 3 2 2 3
end data
manova T1 T2 T3 T4 by HG(1,2)/
  wsfactors=time(4)/
  wsdesign=time/
  print=cellinfo(means)/
  transform/

```

homogeneity(boxm)/
 signif(averf)/
 analysis(repeated)/
 design/
 finish

The means are given below.

Cell Means and Standard Deviations

Variable .. T1					
FACTOR	CODE	Mean	Std. Dev.		N
HG	1	2.500	1.269		10
HG	2	1.700	1.160		10
For entire sample		2.100	1.252		20

--					
Variable .. T2					
FACTOR	CODE	Mean	Std. Dev.		N
HG	1	3.700	1.418		10
HG	2	2.100	1.449		10
For entire sample		2.900	1.619		20

--					
Variable .. T3					
FACTOR	CODE	Mean	Std. Dev.		N
HG	1	3.800	1.229		10
HG	2	2.700	1.252		10
For entire sample		3.250	1.333		20

--					
Variable .. T4					
FACTOR	CODE	Mean	Std. Dev.		N
HG	1	2.700	1.889		10
HG	2	2.200	1.135		10
For entire sample		2.450	1.538		20

The test for homogeneity is not significant, $p = .85$.

Multivariate test for Homogeneity of Dispersion matrices

Boxs M = 7.32300

F WITH (10,1549) DF = .55263, P = .853 (Approx.)

Chi-Square with 10 DF = 5.57361, P = .850 (Approx.)

The between subjects HG effect is significant, $F(1,18) = 5.71$, $p = .028$.

Tests of Between-Subjects Effects.

Tests of Significance for T1 using UNIQUE sums of squares

Source of Variation	SS	DF	MS	F	Sig of F
WITHIN+RESIDUAL	63.05	18	3.50		
HG	20.00	1	20.00	5.71	.028

The epsilon values are given below and are reasonable.

Tests involving 'TIME' Within-Subject Effect.

Mauchly sphericity test, W = .63496
Chi-square approx. = 7.59500 with 5 D. F.
Significance = .180

Greenhouse-Geisser Epsilon = .79888
Huynh-Feldt Epsilon = .98125
Lower-bound Epsilon = .33333

The within subjects effect of Time is significant, $F(3,54) = 3.82$, $p = .015$.

Tests involving 'TIME' Within-Subject Effect.

AVERAGED Tests of Significance for T using UNIQUE sums of squares

Source of Variation	SS	DF	MS	F	Sig of F
WITHIN+RESIDUAL	71.95	54	1.33		
TIME	15.25	3	5.08	3.82	.015
HG BY TIME	3.30	3	1.10	.83	.486

Another Computer Implementation

```
data list/ HG 1 anim 3-4 time 6 Y 8
begin data
1 1 1 3
...
...
...
1 10 4 0
...
...
...
2 10 4 3
end data
manova Y by HG(1,2) anim(1,10) time(1,4)/
      design=HG vs 1, anim WHG=1, time, HG by time/
      omeans=TABLES(HG,time,HG by time)/
      residuals=casewise plot/
finish
```

23.8 Exercises

Hockley (1991) conducted an experiment in which participants studied pairs of words and were later tested for recognition memory for single words (item information) and associations between words (associative information). The principal purpose of this experiment was to determine whether recognition memory is the same or different depending on the nature of the recognition test. Two different types of recognition tests were used. In the yes-no test procedure, a single word. Repeated Measures Types of Designs 431 or a pair of words was presented and participants tried to decide if that word or pair or words appeared in the study list. In the forced-choice test, two single words, or two pairs of words were presented and participants tried to decide which of the two words or pairs had been presented in the study list. The results of the experiment are presented below. Is one test easier than the other or are both types of tests equally difficult?

High Accuracy	TEST			
	YES/NO		F-C	
	TYPE			
	ITEM	ASSOC	ITEM	ASSOC
	S1	1.97	2.36	1.81
S5	2.48	3.22	2.48	3.29
S6	1.43	1.12	1.09	1.19
\bar{x}	1.96	2.23	1.79	2.19
Low Accuracy				
S2	1.32	.95	1.14	.95
S3	.48	.18	.54	.21
S4	.92	.70	.82	.58
\bar{x}	.76	.61	.83	.58